

## Historic, archived document

Do not assume content reflects current scientific knowledge, policies, or practices.



A99.9  
F76254

5



Forest Service

Pacific Northwest  
Research Station

Research Note  
PNW-RN-532  
June 2002



# A Comparison of Two Estimates of Standard Error for a Ratio-of-Means Estimator for a Mapped-Plot Sample Design in Southeast Alaska

Willem W.S. van Hees

## Abstract

Comparisons of estimated standard error for a ratio-of-means (ROM) estimator are presented for forest resource inventories conducted in southeast Alaska between 1995 and 2000. Estimated standard errors for the ROM were generated by using a traditional variance estimator and also approximated by bootstrap methods. Estimates of standard error generated by both traditional and bootstrap methods were similar. Percentage differences between the traditional and bootstrap estimates of standard error for productive forest acres and for gross cubic-foot growth were generally greater than respective differences for non-productive forest acres, net cubic volume, or nonforest acres.

Keywords: Sampling, inventory (forest), error estimation.

## Introduction

Between 1995 and 2000, the Pacific Northwest Research Station (PNW) Forest Inventory and Analysis (FIA) Program in Anchorage, Alaska, conducted a land cover resource inventory of southeast Alaska (fig. 1). Land cover attribute estimates derived from this sample describe such features as area, timber volume, and understory vegetation. Each of these estimates is subject to measurement and sampling error. Measurement error is minimized through training and a program of quality control. Sampling error must be estimated mathematically.

This paper presents a comparison of bootstrap estimation of standard error for a ratio-of-means estimator with a traditional estimation method. Bootstrap estimation techniques can be used when a complex sampling strategy hinders development of unbiased or reliable variance estimators (Schreuder and others 1993.) Bootstrap techniques are resampling methods that treat the sample as if it were the whole population. Repeated samples are taken, with replacement, from the original sample and the statistic of interest is calculated for each sample. The average of the statistic over all the samples is the bootstrap estimator.

This study was conducted to examine briefly whether or not traditional estimates of standard error differed in some meaningful way from the bootstrap approximation. It was felt that the two estimates would produce substantially similar results.

---

Willem W.S. van Hees is a research forester, Forestry Sciences Laboratory, 3301 C. St., Suite 200, Anchorage, AK 99503.

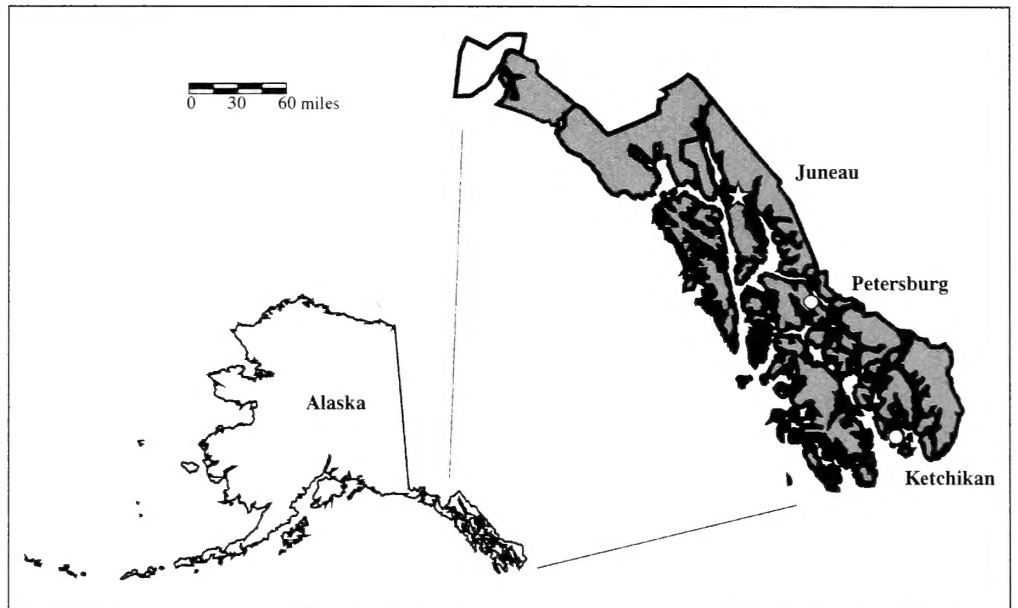


Figure 1—Southeast Alaska inventory region, 1995-2000.

In lieu of extensive simulations, multiple bootstrap estimates of standard error, along with traditional estimates, were made for each of five population parameters in three sampling units. These parameters, for unreserved USDA Forest Service lands, were productive forest-land acres, nonproductive forest-land acres, nonforest acres, net cubic-foot volume, and gross cubic-foot growth in the Chatham, Stikine, and Ketchikan inventory units (van Hees 2001a, 2001b, 2001c).

**Methods**  
**Sample Design**

Land cover of the southeastern panhandle of Alaska (fig. 1) was sampled with a single-phase, unstratified, systematic grid sample (grid spacing was 4.8 kilometers). Ground plots (1 hectare) were established at each grid intersection and were sub-sampled by a cluster of four, 7.3-meter, fixed-radius subplots. Three subplots were sited, one each at 63.4 meters north, southeast, and southwest from the first, centrally located subplot. Each subplot was mapped for land cover (Scott and Bechtold 1995.)

**Ratio-of-Means Estimator**

Estimation of resource attributes such as area and timber volume used ratio-of-means estimators. The ratio-of-means estimator and associated traditional variance estimate are defined as (Cochran 1977):

$$\hat{R} = \frac{\sum_{i=1}^n y_i}{\sum_{i=1}^n x_i} , \tag{1}$$

where

- $y_j$  = the variable of interest on plot  $i$ ,
- $x_j$  = an auxiliary variable (such as area) on plot  $i$  that is correlated with  $y_j$ , and
- $n$  = number of plots selected from the population,

with variance

$$V(\hat{R}) = \frac{1}{n(n-1) \left( \frac{\sum x_i}{n} \right)^2} \left( \sum y_i^2 + \hat{R}^2 \sum x_i^2 - 2\hat{R} \sum y_i x_i \right) . \tag{2}$$

Zarnoch and Bechtold (2000) demonstrate use of this ratio estimator to generate various population estimates from data collected by using a sampling design similar to that described above. Essentially, the ratio estimate defined by equation (1) is multiplied by some total, such as area, to estimate the amount of the total in the condition of interest ( $y_j$ ).

**Table 1—Ratio-of-means (ROM) population estimates, traditional estimate of standard error, and bootstrap estimates of standard error by number of iterations, sampling unit, and population attribute**

Sampling unit/ population attribute	ROM population estimate	No. of plots	Traditional <sup>a</sup> estimate	Estimated standard error		
				500 iterations	1,000 iterations	3,000 iterations
<b>Ketchikan:</b>						
Productive forest, acres	1,405,362	550	58,036	57,410	59,002	58,449
Nonproductive forest, acres	851,262		53,754	55,148	53,482	54,689
Nonforest, acres	449,053		41,962	43,416	41,269	41,961
Net cubic-foot volume	7,293,763,469		483,113,144	504,230,788	462,260,072	479,479,823
Gross cubic-foot growth	63,725,479		3,250,717	3,516,024	3,378,918	3,461,131
<b>Stikine:</b>						
Productive forest, acres	1,085,711	520	60,713	58,648	63,772	57,690
Nonproductive forest, acres	708,385		53,341	51,318	54,061	53,152
Nonforest, acres	1,077,131		60,469	59,758	62,718	58,038
Net cubic-foot volume	5,437,872,712		456,690,704	462,024,168	457,544,019	429,174,042
Gross cubic-foot growth	40,659,537		2,652,600	2,779,894	2,580,381	2,568,629
<b>Chatham:</b>						
Productive forest, acres	1,302,152	979	72,813	69,716	74,637	71,547
Nonproductive forest, acres	709,450		56,555	60,671	56,542	58,272
Nonforest, acres	3,368,739		82,340	83,406	84,722	79,016
Net cubic-foot volume	7,561,291,112		540,235,099	534,119,374	519,603,815	541,402,469
Gross cubic-foot growth	59,953,509		3,803,595	3,856,939	3,654,326	3,786,265
					4,034,644	3,728,864
					3,856,871	3,863,259
					551,209,556	3,929,279
					568,524,342	4,036,473
					536,971,963	568,524,342
					453,095,411	453,530,368
					2,730,045	2,697,633
					468,461,285	470,675,742
					3,414,326	3,446,209

<sup>a</sup> Traditional estimate of standard error is from Cochran (1977).

**Table 2—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, sampling unit, and population attribute**

Sampling unit/ population attribute	Number of bootstrap iterations								
	500			1,000			3,000		
	<i>Percent</i>								
Ketchikan:									
Productive forest, acres	5.42	-1.09	-2.03	1.64			0.71		
Nonproductive forest, acres	5.52	2.53	-0.51	1.64			1.71		
Nonforest, acres	3.79	3.35	-1.68	7.06			0.00		
Net cubic-foot volume	4.19	-4.51	-0.76	-4.24			-3.13	-2.64	-4.30
Gross cubic-foot growth	7.55	3.79	6.08	2.34			4.79	5.67	4.41
Stikine:									
Productive forest, acres	3.18	-3.52	4.80	-5.24			2.89		
Nonproductive forest, acres	1.06	-3.94	1.33	-0.35			3.70		
Nonforest, acres	3.65	-1.19	3.59	-4.19			1.09		
Net cubic-foot volume	1.15	0.19	-1.29	-6.41			-0.79		-0.70
Gross cubic-foot growth	0.29	4.58	-2.80	-3.27			2.84		1.67
Chatham:									
Productive forest, acres	2.09	-4.44	2.44	-3.93	-1.77		-1.52		
Nonproductive forest, acres	-0.88	6.78	-0.02	-0.98	2.95		4.22		
Nonforest, acres	-3.48	1.28	2.81	-3.90	-4.21		0.46		
Net cubic-foot volume	-1.15	-3.97	0.22	3.19	-4.85	-0.95	-0.61	1.99	4.98
Gross cubic-foot growth	1.38	-4.08	-0.46	5.73	-2.00	1.38	1.54	3.20	5.77

### Bootstrap Estimate

Bootstrap estimation of the variance was accomplished by iterative random sampling, with replacement, of the plot list to generate a new plot list. Within each plot, subplots were randomly selected, with replacement, to generate a new subplot list for each plot. The ROM estimates of population totals were then recomputed. The variance of the estimated totals provides an estimate of the sample variance. This process was repeated 500 to 3,000 times. At least four bootstrap estimates, for each of the five population parameters, in all three sampling units, were made by using 500 iterations. Greater numbers of iterations (1,000 and 3,000) were used fewer times and for increasingly fewer population parameters.

Although bootstrap samples are not independent, variance estimates resulting from them are considered conservative approximations (Schreuder and others 1993).

### Results

Bootstrap and traditional estimates of standard errors are presented in table 1. In no case did the bootstrap estimate differ from the traditional estimate by more than 7.55 percent (table 2). Eleven of 92 observations were more than 5 percent different; the remaining 81 bootstrap estimates were less than 4.99 percent different from the ROM estimate. The larger percentage differences (greater than  $\pm 5.5$  percent) are not specific to a particular inventory area or estimated population parameter (figs. 2 through 6).

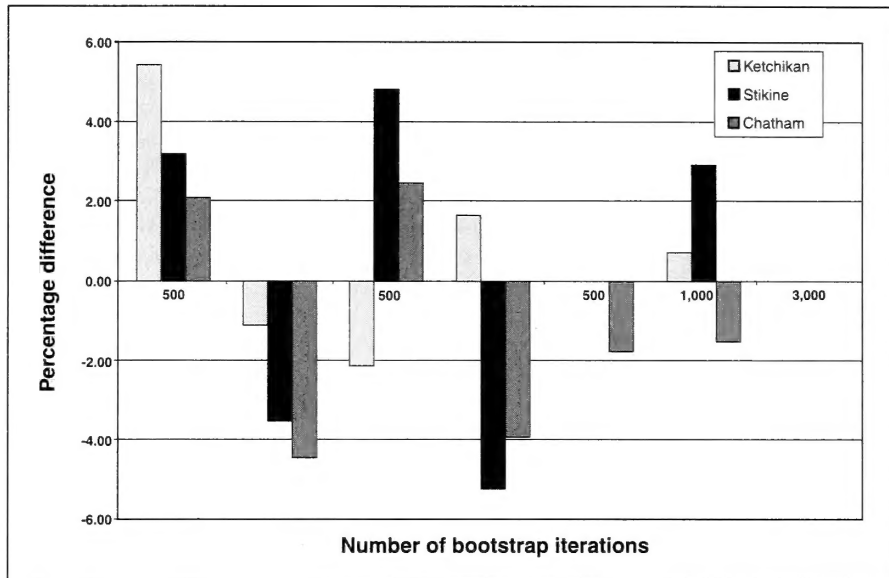


Figure 2—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, for estimates of productive forest-land area within unreserved national forest lands, southeast Alaska.

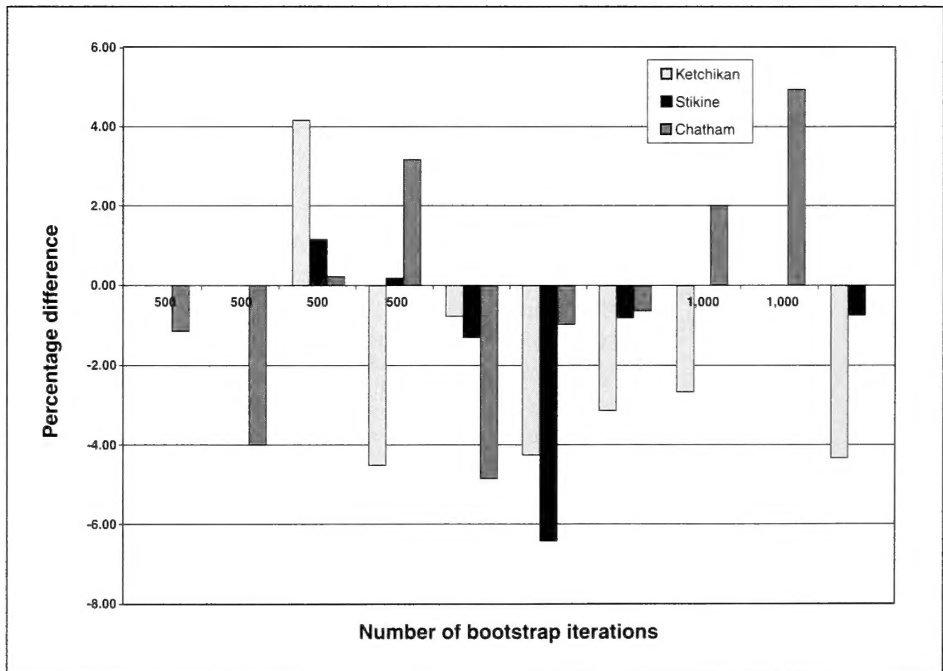


Figure 3—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, for estimates of net cubic-foot volume on productive forest-land area within unreserved national forest lands, southeast Alaska.



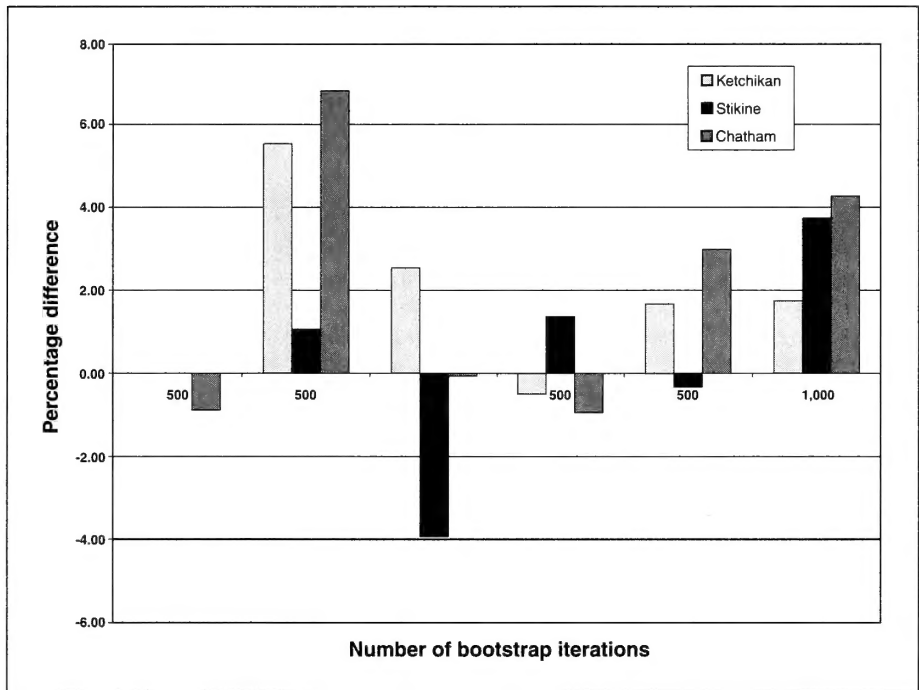


Figure 4—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, for estimates of non-productive forest-land area within unreserved national forest lands, southeast Alaska.

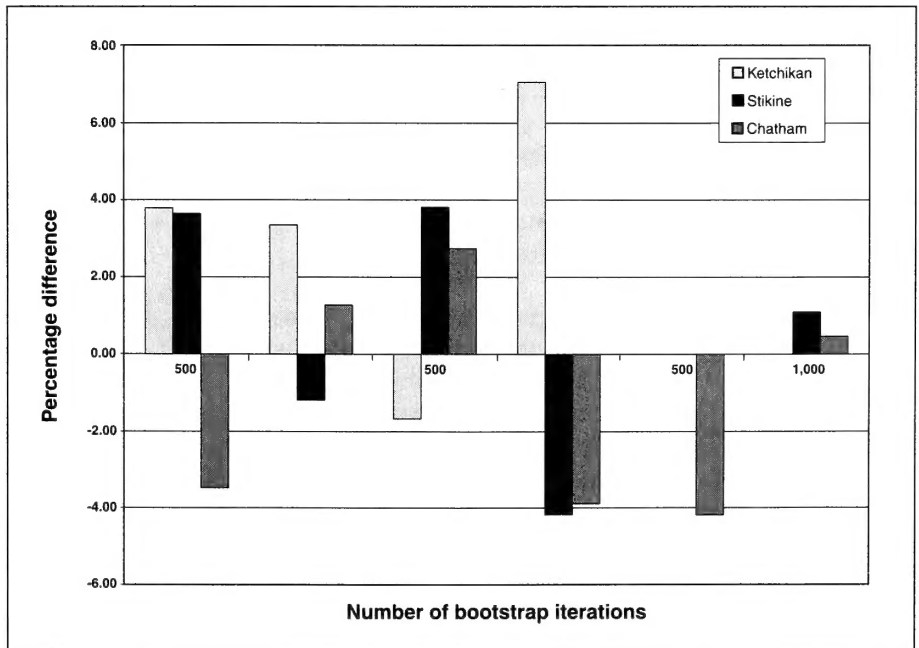


Figure 5—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, for estimates of non-forest-land area within unreserved national forest lands, southeast Alaska.

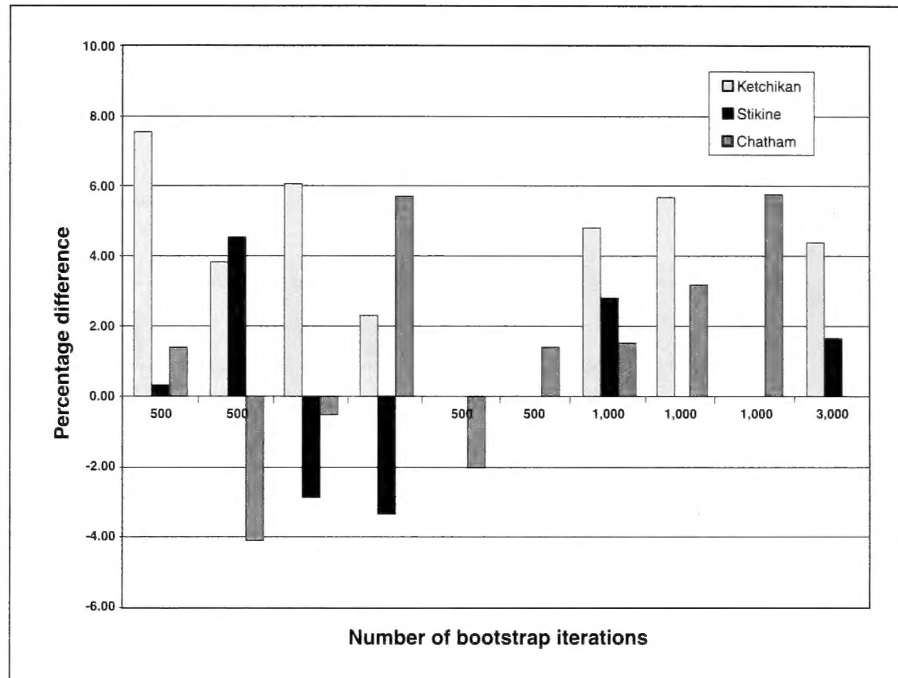


Figure 6—Differences between bootstrap and traditional estimates of standard errors as a percentage of the bootstrap estimate, by number of bootstrap iterations, for estimates of gross cubic-foot growth on productive forest-land area within unreserved national forest lands, southeast Alaska.

## Discussion

For the response variables listed, there is little difference between the two estimates. The largest percentage differences between error estimates are in the Ketchikan unit. The forest resource of the Ketchikan unit is more variable than in the other two units; it is ecologically transitional (approaching the southern extent of the Sitka spruce/western hemlock zone), and logging has been more intensive and extensive resulting in greater variety of age and size classes. Greater variability between bootstrap and traditional estimates may indicate the number of iterations was low for those estimates.

Relative rapidity of data processing is a significant client service consideration. Rapid responses are important when providing clients with inventory results. Although bootstrap estimates can be generated for any of the population estimates FIA produces, the computational resources needed to generate those estimates for each cell in all the output tables typical for FIA reporting exceed those needed for ROM estimates. Use of the ROM estimator reduces production time dramatically.

## Acknowledgments

Kevin Dobelbower, Team Leader for Information Management, Anchorage, Alaska, created and ran necessary computer programs for this study.

## English Equivalents

0.3048 meter = 1 foot  
 0.0283 cubic meter = 1 cubic foot  
 0.4047 hectare = 1 acre  
 1.609 kilometers = 1 mile

## Literature Cited

- Cochran, W.G. 1977.** Sampling techniques. 3<sup>rd</sup> ed. New York: John Wiley and Sons, Inc. 428 p.
- Schreuder, H.T.; Gregoire, T.G.; Wood, G.B. 1993.** Sampling methods for multiresource forest inventory. New York: John Wiley and Sons, Inc. 446 p.
- Scott, C.T.; Bechtold, W.A. 1995.** Techniques and computations for mapping plot clusters that straddle stand boundaries. Forest Science Monograph. 31: 46-61.
- van Hees, W.W.S. 2001a.** Summary estimates of forest resources on unreserved lands of the Chatham inventory unit, Tongass National Forest, southeast Alaska, 1998. Resour. Bull. PNW-RB-234. Portland, OR: U.S. Department of Agriculture, Forest Service, Pacific Northwest Research Station. 12 p.
- van Hees, W.W.S. 2001b.** Summary estimates of forest resources on unreserved lands of the Ketchikan inventory unit, Tongass National Forest, southeast Alaska, 1998. Resour. Bull. PNW-RB-233. Portland, OR: U.S. Department of Agriculture, Forest Service, Pacific Northwest Research Station. 8 p.
- van Hees, W.W.S. 2001c.** Summary estimates of forest resources on unreserved lands of the Stikine inventory unit, Tongass National Forest, southeast Alaska, 1998. Resour. Bull. PNW-RB-232. Portland, OR: U.S. Department of Agriculture, Forest Service, Pacific Northwest Research Station. 8 p.
- Zarnoch, S.J.; Bechtold, W.A. 2000.** Estimating mapped-plot attributes with ratios of means. Canadian Journal of Forest Research. 30: 688-697.



The **Forest Service** of the U.S. Department of Agriculture is dedicated to the principle of multiple use management of the Nation's forest resources for sustained yields of wood, water, forage, wildlife, and recreation. Through forestry research, cooperation with the States and private forest owners, and management of the National Forests and National Grasslands, it strives—as directed by Congress—to provide increasingly greater service to a growing Nation.

The U.S. Department of Agriculture (USDA) prohibits discrimination in all its programs and activities on the basis of race, color, national origin, gender, religion, age, disability, political beliefs, sexual orientation, or marital or family status. (Not all prohibited bases apply to all programs.) Persons with disabilities who require alternative means for communication of program information (Braille, large print, audiotape, etc.) should contact USDA's TARGET Center at (202) 720-2600 (voice and TDD).

To file a complaint of discrimination, write USDA, Director, Office of Civil Rights, Room 326-W, Whitten Building, 14th and Independence Avenue, SW, Washington, DC 20250-9410 or call (202) 720-5964 (voice and TDD). USDA is an equal opportunity provider and employer.

#### **Pacific Northwest Research Station**

<b>Web site</b>	<a href="http://www.fs.fed.us/pnw">http://www.fs.fed.us/pnw</a>
<b>Telephone</b>	(503) 808-2592
<b>Publication requests</b>	(503) 808-2138
<b>FAX</b>	(503) 808-2130
<b>E-mail</b>	<a href="mailto:pnw_pnwpubs@fs.fed.us">pnw_pnwpubs@fs.fed.us</a>
<b>Mailing address</b>	Publications Distribution Pacific Northwest Research Station P.O. Box 3890 Portland, OR 97208-3890

---

U.S. Department of Agriculture  
Pacific Northwest Research Station  
333 S.W. First Avenue  
P.O. Box 3890  
Portland, OR 97208-3890

---

Official Business  
Penalty for Private Use, \$300